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Numerical Integration



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- Numerical Integration
- Monte Carlo Integration









- Numerical Integration
- Monte Carlo Integration
- Quasi Monte Carlo Integration









 $\int_{a}^{b} f(x) dx$





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f(x)dx





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f(x)dx







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$$\int_{a}^{b} f(x) dx$$

Analytic evaluation: accurate and fast









 $\int f(x) dx$

• Numerical evaluations:







$$\int_{a}^{b} f(x) dx$$

- Numerical evaluations:
 - Provide only approximate solutions,







$$\int_{a}^{b} f(x) dx$$

- Numerical evaluations:
 - Provide only approximate solutions,
 - Rate of convergence is important







$$\int_{a}^{b} f(x) dx$$

- Numerical evaluations:
 - Provide only approximate solutions,
 - Rate of convergence is important
 - Often involves evaluations only at selected locations







$$\int_{a}^{b} f(x) dx$$

• Numerical quadrature: designed for 1D integrals





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$$\int_{a}^{b} f(x) dx$$

- Numerical quadrature: designed for 1D integrals
- Cubature/Quadratures: for higher dimensions





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Numerical Integration

Advance Sampling Strategies: June 7, 2018



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• Hybrid methods: First transform the integral analytically for simpler numerical handling



Numerical Integration

- A number of solutions are developed for the numeric solution of integrals
- Most prominent are the Quadrature rules, where the weights w_i and the sample positions x_i are determined in advance

 $\int_{a}^{b} f(x) dx =$



$$= \sum_{i=1}^{N} w_i f(x_i)$$



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- Newton-Cots formula:





(x) x_0 x_{m-1} x_{k}

Composite midpoint formula



• Midpoint rule (1 sample), Trapezoid rule (2 samples), Simpson rule (3 samples)...



- Newton-Cots formula:
 - Midpoint rule (1 sample), Trapezoid rule (2 samples), Simpson rule (3 samples)...
 - Samples are nesting (for powers of 2)
 - Approximates the integral as sum of weighted, equidistant samples









• Gauss quadratures:



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- Gauss quadratures:
 - An n-point Gauss quadrature is constructed to yield exact results for polynomials of degree 2n-1 or less.



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- Gauss quadratures:
 - An n-point Gauss quadrature is constructed to yield exact results for polynomials of degree 2n-1 or less.
 - Extends freedom by allowing choice of sample locations



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 - Extends freedom by allowing choice of sample locations
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Newton-Cots formula*

Gauss quadratures*

smooth integrand that has *r*-continuous derivatives

*Interested students may refer to this link for more details.



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Both approaches achieve convergence of the order $\mathcal{O}(N^{-r})$, given N samples and a

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Numerical Integration: sD case

$$\int_{a}^{b} \dots \int_{a}^{b} f(x_{1}, \dots, x_{s}) dx_{1} \dots dx_{s} = \sum_{i_{1}=1}^{N} \dots \sum_{i_{s}=1}^{N} w_{i_{1}} \dots w_{i_{s}} f(x_{i_{1}}, \dots, x_{i_{s}})$$



• Curse of dimensionality: requires N^s samples for s-dimensional integral



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Numerical Integration: sD case

$$\int_{a}^{b} \dots \int_{a}^{b} f(x_{1}, \dots, x_{s}) dx_{1} \dots dx_{s} = \sum_{i_{1}=1}^{N} \dots \sum_{i_{s}=1}^{N} w_{i_{1}} \dots w_{i_{s}} f(x_{i_{1}}, \dots, x_{i_{s}})$$

- Convergence drops to $\mathcal{O}(N^{-r/s})$



• Curse of dimensionality: requires N^s samples for s-dimensional integral



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Numerical Integration: sD case

$$\int_{a}^{b} \dots \int_{a}^{b} f(x_{1}, \dots, x_{s}) dx_{1} \dots dx_{s} = \sum_{i_{1}=1}^{N} \dots \sum_{i_{s}=1}^{N} w_{i_{1}} \dots w_{i_{s}} f(x_{i_{1}}, \dots, x_{i_{s}})$$

- Convergence drops to $\mathcal{O}(N^{-r/s})$
- Rules must be adapted to non-square domains (typical in rendering)



• Curse of dimensionality: requires N^s samples for s-dimensional integral

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Independent of the dimensions



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- Independent of the dimensions
- Independent of the underlying topology of the domain



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- Independent of the dimensions
- Independent of the underlying topology of the domain



• Variance converges at $O(N^{-1})$ irrespective of the dimensions (N is the sample count)

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- Independent of the dimensions
- Independent of the underlying topology of the domain



• Variance converges at $O(N^{-1})$ irrespective of the dimensions (N is the sample count)

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 $\int_{Q^s} f(x) d\mu_s(x) =$



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 $\int_{Q^s} f(x) d\mu_s(x) = \int_{[0,1)^s} f(x) dx =$



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 $\int_{Q^s} f(x) d\mu_s(x) = \int_{[0,1)^s} f(x) dx = \int_{[0,1)^s} \frac{f(x)}{p(x)} p(x) dx$



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$$\int_{Q^s} f(x) d\mu_s(x) = \int_{[0,1)^s} f(x)$$

 $p(\boldsymbol{x})$: is an arbitrary probability density function over the domain



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 $f(x) = \int_{[0,1)^s} \frac{f(x)}{p(x)} p(x) dx$



 $\int_{[0,1)^s} f(x) dx = \int_{[0,1)^s} \frac{f(x)}{p(x)} p(x) dx$



p(x) : is an arbitrary probability density function over the domain






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p(x) : is an arbitrary probability density function over the domain





 $\int_{[0,1)^s} f(x) dx = \int_{[0,1)^s} f(x) dx$

p(x) : is an arbitrary probability density function over the domain







$$\sum_{0,1)^s} \frac{f(x)}{p(x)} p(x) dx$$

$$\int_{(0,1)^s} \Big(\frac{f(x)}{p(x)}\Big) p(x) dx$$

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 $\int_{[0,1)^s} f(x) dx = \int_{[0,1)^s} f(x) dx$

p(x) : is an arbitrary probability density function over the domain







$$\sum_{0,1)^s} \frac{f(x)}{p(x)} p(x) dx$$

$$\int_{(0,1)^s} \left(\frac{f(x)}{p(x)}\right) p(x) dx$$

$$E[g(x)] = \int_Q g(x)p(x)$$



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 $\int_{[0,1)^s} f(x) dx = \int_{[0,1)^s} f(x) dx$

p(x) : is an arbitrary probability density function over the domain

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$$\sum_{0,1)^s} \frac{f(x)}{p(x)} p(x) dx$$

$$\int_{(0,1)^s} \left(\frac{f(x)}{p(x)}\right) p(x) dx$$

$$= E\left[\frac{f(x)}{p(x)}\right]$$

$$E[g(x)] = \int_Q g(x)p(x)$$







 $\int_{[0,1)^s} f(x)dx = E$

p(x) : is an arbitrary probability density function over the domain



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$$\left[\frac{f(x)}{p(x)}\right]$$

$$E[g(x)] = \int_Q g(x)p(x)$$







 $\int_{[0,1)^s} f(x)dx = E\left[\frac{f(x)}{p(x)}\right]$



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 $\int_{[0,1)^s} f(x)dx = E\left[\frac{f(x)}{p(x)}\right]$

We are interested in the numerical computation of this expected value, leading to The highly important concept of Monte Carlo Estimator



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 $\mathbf{I} = \int_0^1 f(x) dx$





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 $\mathbf{I} = \int_0^1 f(x) dx$









 $\mathbf{I} = \int_0^1 f(x) dx$







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 $\mathbf{I} = \int_0^1 f(x) dx \approx \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}$







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Monte Carlo Estimator in 1D

 $\mathbf{I} = \int_0^1 f(x) dx \approx \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}$

p(x): is the probability density function from which samples are drawn







Monte Carlo Estimator in 2D

 $\mathbf{I} = \int_0^1 f(x) dx \approx \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}$

p(x): is the probability density function from which samples are drawn





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 $\mathbf{I} = \int_0^1 f(x) dx \approx \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}$

p(x): is the probability density function from which samples are drawn







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 $\mathbf{I} = \int_0^1 f(x) dx \approx \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}$

p(x): is the probability density function from which samples are drawn







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$$\mathbf{I}_N = \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}$$

p(x): is the probability density function from which samples are drawn



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Secondary Estimator: $\mathbf{I}_N = \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}$

p(x): is the probability density function from which samples are drawn





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Secondary Estimator:
$$\mathbf{I}_N = \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}$$

 $\mathbf{I}_N = \frac{1}{N} \sum_{i=1}^N \mathbf{I}_1^i$

p(x): is the probability density function from which samples are drawn





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Secondary Estimator:
$$\mathbf{I}_N = \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{p(x_i)}$$

 $\mathbf{I}_N = \frac{1}{N} \sum_{i=1}^N \mathbf{I}_1^i$
Primary Estimator: $\mathbf{I}_1^i = \frac{f(x_i)}{p(x_i)}$

p(x): is the probability density function from which samples are drawn





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Primary Estimator:

$$\mathbf{I}_1^i = \frac{f(x_i)}{p(x_i)}$$



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Primary Estimator:

$$\mathbf{I}_1^i = \frac{f(x_i)}{p(x_i)}$$





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Primary Estimator:

$$\mathbf{I}_1^i = \frac{f(x_i)}{p(x_i)}$$





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Primary Estimator:

$$\mathbf{I}_1^i = \frac{f(x_i)}{p(x_i)}$$





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 $\langle x_i \rangle$ $p(x_i)$





Primary Estimator:

$$\mathbf{I}_1^i = \frac{f(x_i)}{p(x_i)}$$





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Primary Estimator:

$$\mathbf{I}_1^i = \frac{f(x_i)}{p(x_i)}$$







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Primary Estimator:

$$\mathbf{I}_1^i = \frac{f(x_i)}{p(x_i)}$$





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Primary Estimator:

$$\mathbf{I}_1^i = \frac{f(x_i)}{p(x_i)}$$



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$$rac{f(x_{
m R})}{p(x_{
m R})}$$

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Primary Estimator:

$$\mathbf{I}_1^i = \frac{f(x_i)}{p(x_i)}$$





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$$\frac{f(x_i)}{p(x_i)}$$

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Due to the Strong law of large numbers, the arithmetic mean will converge to the expected value with probability 1 given enough samples:

$$\mathbf{I}_N = \frac{1}{N} \sum_{i=1}^N \frac{f(x)}{p(x)}$$



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Due to the Strong law of large numbers, the arithmetic mean will converge to the expected value with probability 1 given enough samples:

$$\operatorname{prob}\left\{\lim_{N\to\infty}\mathbf{I}_N = \frac{1}{N}\sum_{i=1}^N \frac{f(x_i)}{p(x_i)} = \mathbf{E}\left[\frac{f(x)}{p(x)}\right] = \int_Q f(x)dx\right\} = 1$$









Error in Monte Carlo Estimation

$Error = Bias^2 + Variance$

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Monte Carlo estimation is unbiased due to it's "purely" stochastic nature

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Monte Carlo estimation is unbiased due to it's "purely" stochastic nature

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- Monte Carlo estimation is unbiased due to it's "purely" stochastic nature
- rendered images

• We are left with variance, which is visible as stochastic unstructured noise in the

- Monte Carlo estimation is unbiased due to it's "purely" stochastic nature
- rendered images

• We are left with variance, which is visible as stochastic unstructured noise in the

• For biased techniques, it is important to have a consistent solution









- For biased techniques, it is important to have a consistent solution
 - This implies, the bias goes to zero with increase in sample count





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 - Examples: Progressive photon mapping









- For biased techniques, it is important to have a consistent solution
 - This implies, the bias goes to zero with increase in sample count
 - Examples: Progressive photon mapping











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 $\operatorname{Error} = \mathbf{I}_N - \mathbf{I}$

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 $\operatorname{Error} = \mathbf{I}_N - \mathbf{I}$ $\operatorname{Error} = \mathbf{I}_N - \int_O f(x) dx$

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 $\operatorname{Error} = \mathbf{I}_N - \int_Q f(x) dx$

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Bias by definition is the expected error:



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 $\operatorname{Error} = \mathbf{I}_N - \int_Q f(x) dx$

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Bias by definition is the expected error:



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 $\operatorname{Error} = \mathbf{I}_N - \int_O f(x) dx$

Bias = $\mathbf{E}[\text{Error}] = \mathbf{E}\left[\mathbf{I}_N - \int_O f(x)dx\right]$

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Error =

Bias by definition is the expected error:

Bias = E[Err

 $Bias = \mathbf{E} [\mathbf{I}_N]$



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$$= \mathbf{I}_N - \int_Q f(x) dx$$

$$[\operatorname{cor}] = \mathbf{E} \Big[\mathbf{I}_N - \int_Q f(x) dx \Big]$$
$$[\operatorname{V}] - \Big[\int_Q f(x) dx \Big]$$

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$\mathrm{Error} =$

Bias by definition is the expected error:

Bias = E[Err

 $Bias = \mathbf{E} \Big[\mathbf{I}_N \Big]$

 $Bias = \mathbf{E} \left[\mathbf{I}_N \right]$



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$$= \mathbf{I}_N - \int_Q f(x) dx$$

$$\begin{aligned} &\text{ror} \end{bmatrix} = \mathbf{E} \Big[\mathbf{I}_N - \int_Q f(x) dx \Big] \\ &\text{v} \Big] - \Big[\int_Q f(x) dx \Big] \\ &\text{v} \Big] - \int_Q f(x) dx \end{aligned}$$



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Bias = $\mathbf{E}[\mathbf{I}_N] - \int_Q f(x) dx$





$\mathbf{E}[\mathbf{I}_{N}] = \mathbf{E}\left[\frac{1}{N}\sum_{i=1}^{N}\frac{f(x_{i})}{p(x_{i})}\right]$



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Bias = $\mathbf{E}[\mathbf{I}_N] - \int_{\Omega} f(x) dx$





 $\mathbf{E}\left[\mathbf{I}_{N}\right] = \mathbf{E}\left[\frac{1}{N}\sum_{i=1}^{N}\frac{f(x_{i})}{p(x_{i})}\right] = \frac{1}{N}\sum_{i=1}^{N}\mathbf{E}\left[\frac{f(x_{i})}{p(x_{i})}\right]$



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Bias = $\mathbf{E}[\mathbf{I}_N] - \int_O f(x) dx$

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 $\mathbf{E}\Big[\mathbf{I}_{N}\Big] = \mathbf{E}\Big[\frac{1}{N}\sum_{i=1}^{N}\frac{f(x_{i})}{p(x_{i})}\Big] = \frac{1}{N}\sum_{i=1}^{N}\mathbf{E}\Big[\frac{f(x_{i})}{p(x_{i})}\Big] = \frac{1}{N}\sum_{i=1}^{N}\int_{O}\frac{f(x)}{p(x)}p(x)dx$



Bias = $\mathbf{E}[\mathbf{I}_N] - \int_O f(x) dx$





 $\mathbf{E}[\mathbf{I}_{N}] = \mathbf{E}\left[\frac{1}{N}\sum_{i=1}^{N}\frac{f(x_{i})}{p(x_{i})}\right] = \frac{1}{N}\sum_{i=1}^{N}\mathbf{E}\left[\frac{f(x_{i})}{p(x_{i})}\right] = \frac{1}{N}\sum_{i=1}^{N}\int_{Q}\frac{f(x)}{p(x)}p(x)dx$ $=\frac{1}{N}\sum_{i=1}^{N}\int_{O}f(x)dx$



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Bias = $\mathbf{E}[\mathbf{I}_N] - \int_{O} f(x) dx$



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 $\mathbf{E}[\mathbf{I}_{N}] = \mathbf{E}\left[\frac{1}{N}\sum_{i=1}^{N}\frac{f(x_{i})}{p(x_{i})}\right] = \frac{1}{N}\sum_{i=1}^{N}\mathbf{E}\left[\frac{f(x_{i})}{p(x_{i})}\right] = \frac{1}{N}\sum_{i=1}^{N}\int_{Q}\frac{f(x)}{p(x)}p(x)dx$ $=\frac{1}{N}\sum_{i=1}^{N}\int_{Q}f(x)dx$ $\int f(x)dx$



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Bias = $\mathbf{E}[\mathbf{I}_N] - \int_{O} f(x) dx$









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Bias = $\mathbf{E}[\mathbf{I}_N] - \int_O f(x) dx$

 $\mathbf{E}\left[\mathbf{I}_{N}\right] = \int_{O} f(x) dx$

Bias = 0

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For the variance of secondary Monte Carlo Estimator, the following holds:

 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}})$



$$) = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$$









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 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$







$$\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \operatorname{Var}\left(\frac{1}{N}\sum_{i=1}^{N}\frac{f(x_{i})}{p(x_{i})}\right)$$



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 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$







$$\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \operatorname{Var}\left(\frac{1}{N}\sum_{i=1}^{N}\frac{f(x_{i})}{p(x_{i})}\right)$$



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 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$

 $\operatorname{Var}(aX) = a^2 \operatorname{Var}(X)$

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$$\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \operatorname{Var}\left(\frac{1}{N}\sum_{i=1}^{N}\frac{f(x_i)}{p(x_i)}\right) = \frac{1}{N^2}$$



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 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$



 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}})$

$$\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \operatorname{Var}\left(\frac{1}{N}\sum_{i=1}^{N}\frac{f(x_i)}{p(x_i)}\right) = \frac{1}{N^2}$$

 $= \overline{N^2}$

 $\Lambda T2$

⊥ V

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$$\mathbf{J} = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$$
$$\frac{1}{2} \operatorname{Var}\left(\sum_{I=1}^{N} \frac{f(x_i)}{p(x_i)}\right)$$
$$\frac{1}{2} \sum_{I=1}^{N} \operatorname{Var}\left(\frac{f(x_i)}{p(x_i)}\right)$$

$$\operatorname{Var}(aX) = a^2 \operatorname{Var}(X)$$

$$\sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_{1}^{i})$$

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 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}})$

$$\operatorname{Var}(\mathbf{I_N}) = \operatorname{Var}\left(\frac{1}{N}\sum_{i=1}^N \frac{f(x_i)}{p(x_i)}\right) = \frac{1}{N^2}$$

 $= \frac{1}{N^2}$

 $\Lambda T2$

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$$P = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$$

$$\operatorname{Var}\left(\sum_{I=1}^{N} \frac{f(x_i)}{p(x_i)}\right)$$

$$\sum_{I=1}^{N} \operatorname{Var}\left(\frac{f(x_i)}{p(x_i)}\right)$$

$$\sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_{1}^{i})$$

$$\operatorname{Var}(aX) = a^2 \operatorname{Var}$$

Independent samples











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 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$









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 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$

 $\sigma(X) = \sqrt{\operatorname{Var}(X)}$

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 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$

Error = $\sigma(\mathbf{I}_N) = \frac{1}{\sqrt{N^2}} \sqrt{\operatorname{Var}(\mathbf{I}_1^i)}$

 $\sigma(X) = \sqrt{\operatorname{Var}(X)}$





 $\mathrm{Error} = \sigma(\mathbf{I}_N)$



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 $\operatorname{Var}(\mathbf{I}_{\mathbf{N}}) = \frac{1}{N^2} \sum_{i=1}^{N} \operatorname{Var}(\mathbf{I}_1^i)$

$$= \frac{1}{\sqrt{N^2}} \sqrt{\operatorname{Var}(\mathbf{I}_1^i)}$$
$$= \frac{1}{N} \sigma(\mathbf{I}_1^i)$$

 $\sigma(X) = \sqrt{\operatorname{Var}(X)}$









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Convergence rate: MC Estimator









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Convergence rate: MC Estimator









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Convergence rate: MC Estimator









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Convergence rate: MC Estimator











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Convergence rate: MC Estimator













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Inversion methods



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- Inversion methods
- Acceptance-rejection methods



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- Inversion methods
- Acceptance-rejection methods
- Metropolis sampling (later)



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- Inversion methods
- Acceptance-rejection methods
- Metropolis sampling (later)
- Transforming distributions



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Compute the CDF



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• Compute the CDF $P(x) = \int_0^x p(z) dz$



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- Compute the CDF $P(x) = \int_0^x p(z) dz$
- Compute the inverse CDF $P^{-1}(x)$



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- Compute the CDF $P(x) = \int_0^x p(z) dz$
- Compute the inverse CDF $P^{-1}(x)$
- Obtain a uniformly distributed random number $\xi \in [0, 1)$



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- Compute the CDF $P(x) = \int_{0}^{x} p(z)dz$
- Compute the inverse CDF $P^{-1}(x)$
- Obtain a uniformly distributed random number $\xi \in [0, 1)$
- Compute $X_i = P^{-1}(\xi)$



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- Compute the CDF $P(x) = \int_{0}^{x} p(z)dz$
- Compute the inverse CDF $P^{-1}(x)$
- Obtain a uniformly distributed random number $\xi \in [0, 1)$
- Compute $X_i = P^{-1}(\xi)$



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Rendering participating media







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 $p(x) \propto$



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$$p(x) \propto e^{-ax}$$
$$p(x) = ce^{-ax}$$

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 $p(x) \propto e^{-ax}$ $p(x) = ce^{-ax}$

 $\int_0^\infty c e^{-ax} dx = \frac{c}{a} = 1$



 $p(x) \propto$

$$P(x) = \int_0^x c e^{-a}$$



$$\propto e^{-ax}$$
$$= ce^{-ax}$$

 $\int_0^\infty c e^{-ax} dx = \frac{c}{a} = 1$

$^{nx}dx = 1 - e^{-ax} = \xi$





p(x) $\left(\right)$

$$P(x) = \int_0^x c e^{-ax} dx = 1 - e^{-ax} = \xi$$



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$$\propto e^{-ax}$$
$$= ce^{-ax}$$

$$\int_0^\infty c e^{-ax} dx = \frac{c}{a}$$

$$P^{-1}(x) = \frac{\ln(1 - x)}{a}$$









p(x)p(x)

$$P(x) = \int_0^x c e^{-ax} dx = 1 - e^{-ax} = \xi$$
$$P^{-1}(x) = \frac{\ln(1-\xi)}{a}$$



$$\propto e^{-ax}$$
$$= ce^{-ax}$$

$$\int_0^\infty c e^{-ax} dx = \frac{c}{a}$$

$$P^{-1}(x) = \frac{\ln(1 - x)}{a}$$



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p(x)p(x)

$$P(x) = \int_0^x ce^{-ax} dx = 1 - e^{-ax} = \xi$$
$$P^{-1}(x) = \frac{\ln(1-\xi)}{a}$$
$$P^{-1}(x) = \frac{\ln(\xi)}{a}$$



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Many samples are wasted







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- Many samples are wasted
- Very costly







- Many samples are wasted
- Very costly
- Not possible for arbitrary domains



with a function f.



 General question: which distributions results when we transform samples from an arbitrary distributions to some other distribution

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with a function f.

 $X_i \sim p_x(x)$



 General question: which distributions results when we transform samples from an arbitrary distributions to some other distribution

Realistic Image Synthesis SS2018





with a function f.

 $X_i \sim p_x(x)$

$$Y_i = y(X_i)$$



 General question: which distributions results when we transform samples from an arbitrary distributions to some other distribution

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with a function f.

$$X_i \sim p_x(x)$$

$$Y_i = y(X_i)$$

What is the distribution of Y_i ?



 General question: which distributions results when we transform samples from an arbitrary distributions to some other distribution

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• The function y(x) must be a one-to-one transformation









- The function y(x) must be a one-to-one transformation
 - It's derivative must either be strictly > 0 or strictly < 0









- The function y(x) must be a one-to-one transformation
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$\operatorname{prob}\{Y \le y(x$



$$x)\} = \operatorname{prob}\{X \le x\}$$







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 $\operatorname{prob}\{Y \le y(x)\} = \operatorname{prob}\{X \le x\}$

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 $\operatorname{prob}\{Y \le y(x)\} = \operatorname{prob}\{X \le x\}$

 $P_y(y) = P_y(y(x)) = P_x(x)$

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- $\operatorname{prob}\{Y \le y(x)\} = \operatorname{prob}\{X \le x\}$
 - $P_{y}(y) = P_{y}(y(x)) = P_{x}(x)$
- This relationship between CDFs directly leads to the relationship between their PDFs:







- $\operatorname{prob}\{Y \le y(x)\} = \operatorname{prob}\{X \le x\}$
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 $p_y(y)$



This relationship between CDFs directly leads to the relationship between their PDFs:

$$\frac{dy}{dx} = p_x(x)$$





- $p_y(y)$
- $p_y(y) =$



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This relationship between CDFs directly leads to the relationship between their PDFs:

$$\frac{dy}{dx} = p_x(x)$$
$$= \left(\frac{dy}{dx}\right)^{-1} p_x(x)$$




$p_y(y) =$



$$= \left(\frac{dy}{dx}\right)^{-1} p_x(x)$$

In general, the derivative is strictly positive or negative, and the relationship between the densities is:

$$p_y(y) = \left| \frac{dy}{dx} \right|^{-1} p_x(x)$$



How can we use this formula?



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 $p_y(y) = \left| \frac{dy}{dx} \right|^{-1} p_x(x)$





How can we use this formula?

$$p_x(x) = 2x \qquad x \in [0, 1]$$
$$Y = \sin X$$



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 $p_y(y) = \left| \frac{dy}{dx} \right|^{-1} p_x(x)$

$$p_y(y) = \frac{p_x(x)}{|\cos x|} = \frac{2x}{\cos x} = \frac{2\arcsin y}{\sqrt{1-y^2}}$$



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given transformation



• Usually we have some PDF that we want to sample from, not a

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- given transformation
- like to compute $Y \sim p_y(y)$



• Usually we have some PDF that we want to sample from, not a

• For example, we might have given: $X \sim p_x(x)$ and we would



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 $y(x) = P_u^{-1}(P_x(x))$



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- Usually we have some PDF that we want to sample from, not a given transformation
- For example, we might have given: $X \sim p_x(x)$ and we would like to compute $Y \sim p_y(y)$

$$P_y(y) = P_x(x)$$

• This is a generalization of the inversion method.



$$y(x) = P_y^{-1}(P_x(x))$$







Transformation in Multiple dimensions

- Suppose we have an s-dimensional X with density function p_X
- Now let Y = T(X) where T is a bijection.





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Transformation in Multiple dimensions

- Suppose we have an s-dimensional X with density function p_X
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 $p_y(y) = p_y(y)$



$$(T(x)) = \frac{p_x(x)}{|J_T(x)|}$$



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Transformation in Multiple dimensions

- Suppose we have an s-dimensional X with density function p_X
- Now let Y = T(X) where T is a bijection.

 $p_y(y) = p_y(y)$

$$J_T(x) = \begin{pmatrix} \partial T_1 / \partial x_1 \\ \vdots \\ \partial T_n / \partial x_1 \end{pmatrix}$$



$$\begin{aligned} (T(x)) &= \frac{p_x(x)}{|J_T(x)|} \\ \cdots & \partial T_1 / \partial x_n \\ \vdots \\ \cdots & \partial T_n / \partial x_n \end{aligned} \end{aligned}$$

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Suppose we draw samples from some density $p(r, \theta)$



- $x = r\cos\theta$
- $y = r\sin\theta$

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- $x = r\cos\theta$
- $y = r\sin\theta$
- Suppose we draw samples from some density $p(r, \theta)$
 - What is the corresponding density p(x, y)?







$$J_T = \begin{pmatrix} \frac{\partial x}{\partial r} & \frac{\partial x}{\partial \theta} \\ \frac{\partial y}{\partial r} & \frac{\partial y}{\partial \theta} \end{pmatrix} = \begin{pmatrix} \cos \theta & -r \sin \theta \\ \sin \theta & r \cos \theta \end{pmatrix}$$



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$$p(x,y) = p(r,\theta)/J_T$$





$$J_T = \begin{pmatrix} \frac{\partial x}{\partial r} & \frac{\partial x}{\partial \theta} \\ \frac{\partial y}{\partial r} & \frac{\partial y}{\partial \theta} \end{pmatrix} = \begin{pmatrix} \cos \theta & -r \sin \theta \\ \sin \theta & r \cos \theta \end{pmatrix}$$



- $x = r\cos\theta$
- $y = r\sin\theta$
- Suppose we draw samples from some density $p(r, \theta)$
 - What is the corresponding density p(x, y)?

$$p(x, y) = p(r, \theta) / J_T$$
$$p(x, y) = p(r, \theta) / r$$





 $z = r \cos \theta$,



Spherical Coordinates

 $x = r \sin \theta \cos \phi$ $y = r \sin \theta \sin \phi$

- $|J_T| = r^2 \sin \theta$
- $p(r, \theta, \phi) = r^2 \sin \theta \ p(x, y, z)$





Spherical coordinates

$$|J_T| = r^2 \sin \theta$$

 $p(r, \theta, \phi) = r^2 \sin \theta \ p(x, y, z)$

 $p(\theta, \phi) d\theta$ $p(\theta$



Spherical Coordinates

 $x = r \sin \theta \cos \phi$ $y = r \sin \theta \sin \phi$ $z = r \cos \theta$,

 $d\omega = \sin\theta \, \mathrm{d}\theta \, \mathrm{d}\phi$

$$Pr\left\{\omega\in\Omega\right\}=\int_{\Omega}p(\omega)\,\mathrm{d}\omega$$

$$\theta d\phi = p(\omega) d\omega$$

, $\phi) = \sin \theta \ p(\omega)$

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Here, the task is to choose a direction on the hemisphere uniformly w.r.t. solid angle. Using the fact that, PDF must integrate to one over its domain:







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$$\int_{\mathcal{H}^2} p(\omega) \, \mathrm{d}\omega = 1 \Rightarrow c \, \int_{\mathcal{H}^2} \mathrm{d}\omega = 1 \Rightarrow c = \frac{1}{2\pi}$$





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Conditional density function:





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$$p(\phi|\theta) = \frac{p(\theta, \phi)}{p(\theta)} = \frac{1}{2\pi}$$

Marginal density function: μ

$$\int_{\mathcal{H}^2} p(\omega) \, d\omega = 1 \Rightarrow c \int_{\mathcal{H}^2} d\omega = 1 \Rightarrow c = \frac{1}{2\pi} \qquad p(\omega) = 1/(2\pi)$$

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Conditional density function:



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Corresponding CDFs:



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Corresponding CDFs:

 $P(\theta) = \int_0^\theta \sin \theta' \, \mathrm{d}\theta' = 1 - \cos \theta$ $P(\phi|\theta) = \int_0^\phi \frac{1}{2\pi}$



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$$\mathrm{d}\phi' = \frac{\phi}{2\pi}.$$







Uniformly sampling a hemisphere $P(\theta) = \int_{0}^{\theta} \sin \theta' \, \mathrm{d}\theta' = 1 - \cos \theta$ Corresponding CDFs: $P(\phi|\theta) = \int_0^{\phi} \frac{1}{2\pi}$

Inverting these functions is straightforward, and here we can safely write:



$$\mathrm{d}\phi' = \frac{\phi}{2\pi}.$$



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Uniformly sampling

$$P(\theta) = \int_{0}^{\theta} \sin \theta' \, d\theta' = 1 - \frac{1}{2\pi} \, d\theta' = \frac{1}{2\pi} \, d\theta' =$$

Inverting these functions is straightforward, and here we can safely write:

$$\theta = \cos^{-1} \xi_1$$
$$\phi = 2\pi \xi_2.$$



a hemisphere

 $-\cos\theta$



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Uniformly sampling

$$P(\theta) = \int_{0}^{\theta} \sin \theta' \, d\theta' = 1 - \frac{1}{2\pi} \, d\theta' = \frac{1}{2\pi} \, d\theta' =$$

Inverting these functions is straightforward, and here we can safely write:

$$\theta = \cos^{-1} \xi_1$$
$$\phi = 2\pi \xi_2.$$



a hemisphere

 $-\cos\theta$

$$x = \sin \theta \cos \phi = \cos \left(2\pi \xi_2\right) \sqrt{1 - \xi_1^2}$$
$$y = \sin \theta \sin \phi = \sin \left(2\pi \xi_2\right) \sqrt{1 - \xi_1^2}$$
$$z = \cos \theta = \xi_1.$$





Uniformly sampling a disk





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Correct PDF ???




Uniformly sampling a disk

Marginal density function:

Conditional density function:



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- $p(x, y) = 1/\pi$
- $p(r, \theta) = r/\pi$

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Uniformly sampling a disk $p(x, y) = 1/\pi$ $p(r, \theta) = r/\pi$

 $p(r) = \int_{0}^{2\pi} p(r, \theta) \, \mathrm{d}\theta = 2r$ Marginal density function: Conditional density function: $p(\theta|r) = \frac{p(r, \theta)}{p(r)} = \frac{1}{2\pi}$.





Uniformly sampling a disk $p(x, y) = 1/\pi$ $p(r, \theta) = r/\pi$

 $p(r) = \int_{0}^{2\pi} p(r, \theta) \, \mathrm{d}\theta = 2r$ Marginal density function: Conditional density function: $p(\theta|r) = \frac{p(r, \theta)}{p(r)} = \frac{1}{2\pi}$.



- $r = \sqrt{\xi_1}$
- $\theta = 2\pi\xi_2.$



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Importance Sampling



66

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- Importance Sampling
- Multiple Importance Sampling



66

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- Importance Sampling
- Multiple Importance Sampling
- Control Variates



66

Realistic Image Synthesis SS2018





- Importance Sampling
- Multiple Importance Sampling
- Control Variates
- Stratified Sampling



66

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Importance Sampling doesn't always reduce variance.



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 $\mathbf{I}_N = \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$





 $\mathbf{I}_N =$

- Importance Sampling doesn't always reduce variance.



$$\frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$$

• The pdf $p(\vec{x})$ must be carefully chosen to gain improvements

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 $\mathbf{I}_N = \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$







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 $\mathbf{I}_N = \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$

 $p(\vec{x}) \propto f(\vec{x})$



 $\mathbf{I}_N = \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$ $p(\vec{x}) \propto f(\vec{x})$ $p(\vec{x}) = cf(\vec{x})$





 $\mathbf{I}_N = \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$ $p(\vec{x}) \propto f(\vec{x})$ $p(\vec{x}) = cf(\vec{x})$



 $\int_{-\infty}^{\infty} p(\vec{x}) d\vec{x} = 1$



 $\mathbf{I}_N = \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$ $p(\vec{x}) \propto f(\vec{x})$

 $p(\vec{x}) = cf(\vec{x})$



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 $\int_{-\infty}^{\infty} p(\vec{x}) d\vec{x} = 1$

 $\int_{-\infty}^{\infty} cf(\vec{x})d\vec{x} = 1$





 $\mathbf{I}_N =$

 $p(\bar{x}$

 $p(\vec{x})$

c = -



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$$= \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$$
$$\vec{x} \propto f(\vec{x})$$
$$\vec{x} = cf(\vec{x})$$
$$\frac{1}{\int_{-\infty}^{\infty} f(\vec{x})} d\vec{x}$$

$$\int_{-\infty}^{\infty} p(\vec{x}) d\vec{x} = 1$$

 $\int_{-\infty}^{\infty} cf(\vec{x})d\vec{x} = 1$



 $\mathbf{I}_N =$

 $p(\vec{x})$

 $p(\vec{x})$

 $c = \frac{1}{\int c}$

this seems like a no-op since the PDF computation requires the integral of the function that we are interested in estimating.



$$= \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$$

$$\vec{x} \propto f(\vec{x})$$

$$\int_{-\infty}^{\infty} p(\vec{x}) d\vec{x} = 1$$

$$\int_{-\infty}^{\infty} cf(\vec{x}) d\vec{x} = 1$$

$$\int_{-\infty}^{\infty} cf(\vec{x}) d\vec{x} = 1$$





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 $\mathbf{I}_N = \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$ $p(\vec{x}) = \frac{f(\vec{x})}{\int_{-\infty}^{\infty} f(\vec{x})} d\vec{x}$



 $p(\vec{x}) =$

 $\mathbf{I}_N =$



 $\mathbf{I}_N = \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$

$$\frac{f(\vec{x})}{\int_{-\infty}^{\infty} f(\vec{x})} d\vec{x}$$

$$\int_{-\infty}^{\infty} f(\vec{x}) d\vec{x}$$





 $\mathbf{I}_N =$

 $p(\vec{x}) =$

 $\mathbf{I}_N =$

• However, this is a very special case that we are encountering here.



$$= \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$$
$$= \frac{f(\vec{x})}{\int_{-\infty}^{\infty} f(\vec{x})} d\vec{x}$$
$$\int_{-\infty}^{\infty} f(\vec{x}) d\vec{x}$$



 $\mathbf{I}_N =$

 $p(\vec{x}) =$

 $\mathbf{I}_N =$

- However, this is a very special case that we are encountering here.



$$= \frac{1}{N} \frac{f(\vec{x_i})}{p(\vec{x_i})}$$
$$= \frac{f(\vec{x})}{\int_{-\infty}^{\infty} f(\vec{x})} d\vec{x}$$
$$\int_{-\infty}^{\infty} f(\vec{x}) d\vec{x}$$

This is referred to as Perfect Importance Sampling, for which the variance is zero.









 $f(ec{x})$

Examples of perfect importance sampling for which the variance is zero



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 $f(\vec{x})$

Examples of perfect importance sampling for which the variance is zero



Realistic Image Synthesis SS2018







 $f(\vec{x})$

Examples of perfect importance sampling for which the variance is zero





 $g(ec{x})$







 $f(\vec{x})$

Examples of perfect importance sampling for which the variance is zero





$$g(\vec{x})$$





Scattering equation:





Image from PBRT 2016



Scattering equation:

$$L_{o}(\mathbf{p}, \omega_{o}) = \int_{\mathbb{S}^{2}} f(\mathbf{p}, \omega_{o}, \omega_{i}) L_{i}(\mathbf{p}, \omega_{i}) |\cos \theta_{i}| d\omega_{i}$$
$$\approx \frac{1}{N} \sum_{j=1}^{N} \frac{f(\mathbf{p}, \omega_{o}, \omega_{j}) L_{i}(\mathbf{p}, \omega_{j}) |\cos \theta_{j}|}{p(\omega_{j})}$$

Cosine weighted spherical/hemispherical sampling









Scattering equation:

$$L_{o}(\mathbf{p}, \omega_{o}) = \int_{\mathbb{S}^{2}} f(\mathbf{p}, \omega_{o}, \omega_{i}) L_{i}(\mathbf{p}, \omega_{i}) |\cos \theta_{i}| d\omega_{i}$$
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Cosine weighted spherical/hemispherical sampling





$$p(\omega) \propto \cos \theta$$





Reference image N = 1024 spp



BSDF importance sampling N = 4 spp

Light importance sampling N = 4 spp







Reference image N = 1024 spp



BSDF importance sampling

Light importance sampling

N = 4 spp

- N = 4 spp
- BSDF sampling is better in some regions





Reference image N = 1024 spp

BSDF importance sampling N = 4 spp



Light importance sampling

N = 4 spp

Light sampling is better in other regions











Reference image

Can we combine the benefits of different PDFs ?



BSDF importance sampling

Light importance sampling







BSDF importance sampling

Light importance sampling



Can we combine the benefits of different PDFs?









BSDF importance sampling

Light importance sampling



Can we combine the benefits of different PDFs ? Yes!









BSDF importance sampling

Light importance sampling

Can we combine the benefits of different PDFs ? Yes!



Multiple Importance Sampling







Variance reduction: Multiple Importance sampling

Multiple Importance Sampling



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 $I_{N} = \frac{1}{N} \sum_{i=1}^{N} \frac{f(x)g(x)}{p(x)}$




Variance reduction: Multiple Importance sampling

Multiple Importance Sampling

 $p(x) \propto ???$



Realistic Image Synthesis SS2018

 $I_{N} = \frac{1}{N} \sum_{i=1}^{N} \frac{f(x)g(x)}{p(x)}$



Variance reduction: Multiple Importance sampling

Multiple Importance Sampling

 $I_N = -\frac{1}{\Lambda}$

 $p(x) \propto ???$

$$\mathbf{I}_N = \frac{1}{n_f} \sum_{i=1}^{n_f} \frac{f(X_i)g(X_i)w_f(X_i)}{p_f(X_i)} + \frac{1}{n_g} \sum_{j=1}^{n_g} \frac{f(Y_j)g(Y_j)w_g(Y_j)}{p_g(Y_j)}$$



$$\frac{1}{N} \sum_{i=1}^{N} \frac{f(x)g(x)}{p(x)}$$





Variance reduction: Multiple Importance sampling

Multiple Importance Sampling

$$\mathbf{I}_N = \frac{1}{n_f} \sum_{i=1}^{n_f} \frac{f(X_i)g(X_i)w_f(X_i)}{p_f(X_i)} + \frac{1}{n_g} \sum_{j=1}^{n_g} \frac{f(Y_j)g(Y_j)w_g(Y_j)}{p_g(Y_j)}$$

Balance heuristic: $w_s(x) =$

Power heuristic: $w_s(x) =$



$$= \frac{n_s p_s(x)}{\sum_i n_i p_i(x)}$$

$$= \frac{(n_s p_s(x))^{\beta}}{\sum_i (n_i p_i(x))^{\beta}}$$

$$\beta = 2$$

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integrand is sought



• To reduce variance, an easily evaluated approximation to the





- integrand is sought
- make use of correlated points in the sampling



• To reduce variance, an easily evaluated approximation to the

• Instead sampling all points independently, control variates



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- integrand is sought
- make use of correlated points in the sampling
- The mathematical basis of control variates is the linearity integral under study.



• To reduce variance, an easily evaluated approximation to the

• Instead sampling all points independently, control variates

property of the Lebesgue integral, i.e., one try to find an analytically Lebesgue-integrable function g that is similar to the



 $\int_Q f(x) dx =$



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 $\int_Q f(x)dx = \int_Q g(x)dx + \int_Q (f(x) - g(x))dx$



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 $\int_{Q} f(x) dx = \int_{Q} g(x)$



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$$\begin{aligned} x &= \int_Q g(x)dx + \int_Q (f(x) - g(x))dx \\ &= \int_Q g(x)dx + \int_Q \frac{(f(x) - g(x))}{p(x)} p(x)dx \end{aligned}$$



 $\int_{Q} f(x) dx = \int_{Q} g(x)$

 $= \int_{Q} g(x)$

 $= \int_{Q} g(x)$



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$$\begin{aligned} f(x)dx &+ \int_{Q} (f(x) - g(x))dx \\ f(x)dx &+ \int_{Q} \frac{(f(x) - g(x))}{p(x)} p(x)dx \\ f(x)dx &+ \mathbf{E} \left[\frac{(f(x) - g(x))}{p(x)} \right] \end{aligned}$$



 $\int_{Q} f(x) dx = \int_{Q} g(x)$

Since we don't know the analytic integral solution of f(x)the corresponding estimator can be written as:



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$$(x)dx + \mathbf{E}\left[\frac{(f(x) - g(x))}{p(x)}\right]$$





 $\int_{\Omega} f(x) dx = \int_{\Omega} g(x)$

Since we don't know the analytic integral solution of f(x)the corresponding estimator can be written as:

 $\mathbf{I}_N^{CV} = \int_O g(x) dx +$



$$(x)dx + \mathbf{E}\left[\frac{(f(x) - g(x))}{p(x)}\right]$$

$$\frac{1}{N} \sum_{i=1}^{N} \left[\frac{(f(x_i) - g(x_i))}{p(x_i)} \right]$$



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 $\mathbf{I}_{N}^{CV} = \int_{Q} g(x)dx + \frac{1}{N} \sum_{i=1}^{N} \left| \frac{(f(x_{i}) - g(x_{i}))}{p(x_{i})} \right|$



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$$\mathbf{I}_{N}^{CV} = \int_{Q} g(x)dx + \frac{1}{N} \sum_{i=1}^{N} \left[\frac{(f(x_{i}) - g(x_{i}))}{p(x_{i})} \right]$$

The integral on the right hand side can be evaluated exactly, where as the variance of the estimator is given by:







$$\mathbf{I}_{N}^{CV} = \int_{Q} g(x)dx + \frac{1}{N} \sum_{i=1}^{N} \left[\frac{(f(x_{i}) - g(x_{i}))}{p(x_{i})} \right]$$

The integral on the right hand side can be evaluated exactly, where as the variance of the estimator is given by:

$$\operatorname{Var}(\mathbf{I}_{N}^{CV}) = \frac{1}{N^{2}} \sum_{i=1}^{N} \operatorname{Var}\left(\frac{(f(x_{i}) - g(x_{i}))}{p(x_{i})}\right)$$







$$\mathbf{I}_{N}^{CV} = \int_{Q} g(x)dx + \frac{1}{N} \sum_{i=1}^{N} \left[\frac{(f(x_{i}) - g(x_{i}))}{p(x_{i})} \right]$$

The integral on the right hand side can be evaluated exactly, where as the variance of the estimator is given by:

$$\operatorname{Var}(\mathbf{I}_{N}^{CV}) = \frac{1}{N^{2}} \sum_{i=1}^{N} \operatorname{Var}\left(\frac{(f(x_{i}) - g(x_{i}))}{p(x_{i})}\right)$$

Variance can be reduced if:







$$\mathbf{I}_{N}^{CV} = \int_{Q} g(x)dx + \frac{1}{N} \sum_{i=1}^{N} \left[\frac{(f(x_{i}) - g(x_{i}))}{p(x_{i})} \right]$$

The integral on the right hand side can be evaluated exactly, where as the variance of the estimator is given by:

$$\operatorname{Var}(\mathbf{I}_{N}^{CV}) = \frac{1}{N^{2}} \sum_{i=1}^{N} \operatorname{Var}\left(\frac{(f(x_{i}) - g(x_{i}))}{p(x_{i})}\right)$$

Variance can be reduced if:

$$\operatorname{Var}\left(\frac{(f(x_i) - g(x_i))}{p(x_i)}\right) < \operatorname{Var}\left(\frac{f(x_i)}{p(x_i)}\right)$$









Jittered Sampling

Latin Hypercube Sampling



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Random 2D







0

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0

87

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Random Samples













Random Samples



Stratified sampling suffers from the curse of dimensionality



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Jittered Samples

N = 64 spp



Jittered Sampling

Latin Hypercube Sampling



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Initialize



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Shuffle rows

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Shuffle rows



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Realistic Image Synthesis SS2018



Shuffle columns





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Shuffle columns





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Variants of stratified sampling



Figure 2.25: Stratification of I² with Voronoi diagrams. (a) 64-element Hammersley point set; (b) Voronoi diagram implied through (a); (c) 64-element hexagonal grid; (d) Voronoi diagram implied through (c).





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Monte Carlo integration suffers, apart from the slow convergence rate, from the are possible



disadvantages that only probabilistic statements on convergence and error boundaries

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Monte Carlo integration suffers, apart from the slow convergence rate, from the are possible



disadvantages that only probabilistic statements on convergence and error boundaries

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- Monte Carlo integration suffers, apart from the slow convergence rate, from the are possible
- random samples



disadvantages that only probabilistic statements on convergence and error boundaries

The success of any Monte Carlo procedure stands or falls with the quality of these





- Monte Carlo integration suffers, apart from the slow convergence rate, from the are possible
- random samples



disadvantages that only probabilistic statements on convergence and error boundaries

The success of any Monte Carlo procedure stands or falls with the quality of these





- Monte Carlo integration suffers, apart from the slow convergence rate, from the are possible
- random samples
- there are no samples at all, which can increases the error



disadvantages that only probabilistic statements on convergence and error boundaries

The success of any Monte Carlo procedure stands or falls with the quality of these

• If the distribution of the sample points is not uniform then there are large regions where

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- Monte Carlo integration suffers, apart from the slow convergence rate, from the are possible
- random samples
- there are no samples at all, which can increases the error



disadvantages that only probabilistic statements on convergence and error boundaries

The success of any Monte Carlo procedure stands or falls with the quality of these

• If the distribution of the sample points is not uniform then there are large regions where

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- Monte Carlo integration suffers, apart from the slow convergence rate, from the are possible
- random samples
- there are no samples at all, which can increases the error
- many locations if samples are clumped



disadvantages that only probabilistic statements on convergence and error boundaries

The success of any Monte Carlo procedure stands or falls with the quality of these

• If the distribution of the sample points is not uniform then there are large regions where

• Closely related to this is the fact that a smooth function is evaluated at unnecessary



Deterministic generation of samples, while making sure uniform distributions



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Deterministic generation of samples, while making sure uniform distributions



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Deterministic generation of samples, while making sure uniform distributions



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- Deterministic generation of samples, while making sure uniform distributions
- Based on number-theoretic approaches



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- Deterministic generation of samples, while making sure uniform distributions
- Based on number-theoretic approaches



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- Deterministic generation of samples, while making sure uniform distributions
- Based on number-theoretic approaches



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- Deterministic generation of samples, while making sure uniform distributions
- Based on number-theoretic approaches
- Samples with good uniform properties can be generated in very high dimensions.



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- Deterministic generation of samples, while making sure uniform distributions
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- Deterministic generation of samples, while making sure uniform distributions
- Based on number-theoretic approaches
- Samples with good uniform properties can be generated in very high dimensions.



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- Deterministic generation of samples, while making sure uniform distributions
- Based on number-theoretic approaches
- Samples with good uniform properties can be generated in very high dimensions.
- Sample generation is pretty fast: (almost) no pre-processing



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• Low discrepancy sequences



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- Low discrepancy sequences
 - Halton and Hammerslay sequences



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- Low discrepancy sequences
 - Halton and Hammerslay sequences
 - Scrambled sequences



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- Low discrepancy sequences
 - Halton and Hammerslay sequences
 - Scrambled sequences
- Discrepancy



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- Low discrepancy sequences
 - Halton and Hammerslay sequences
 - Scrambled sequences
- Discrepancy
- Koksma-Hlawka Inequality (later)



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distribution



• The concept of discrepancy can be viewed as a quantitative measure for the deviation of a given point set from a uniform





distribution





• The concept of discrepancy can be viewed as a quantitative measure for the deviation of a given point set from a uniform





distribution





• The concept of discrepancy can be viewed as a quantitative measure for the deviation of a given point set from a uniform

Area of the blue box:





distribution





• The concept of discrepancy can be viewed as a quantitative measure for the deviation of a given point set from a uniform

Area of the blue box: 0.09





distribution





• The concept of discrepancy can be viewed as a quantitative measure for the deviation of a given point set from a uniform

> Area of the blue box: 0.09 Area associated to each sample: 0.25

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distribution





• The concept of discrepancy can be viewed as a quantitative measure for the deviation of a given point set from a uniform

> Area of the blue box: 0.09 Area associated to each sample: 0.25 Discrepancy:





distribution





• The concept of discrepancy can be viewed as a quantitative measure for the deviation of a given point set from a uniform

> Area of the blue box: 0.09 Area associated to each sample: 0.25 Discrepancy: 0.25 - 0.09 = 0.16





Any integer can be represented in the form:

$$n = \sum_{i=1}^{\infty} d_i b^{i-1}$$





Techniques based on a construction called as radical inverse

n	Binary	$\Phi_b(n)$
1	1	
2	01	
3	11	
4	001	
5	101	



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Any integer can be represented in the form:

$$n = \sum_{i=1}^{\infty} d_i b^{i-1}$$

Radical inverse:

$$\Phi_b(n) = 0.d_1d_2...d_m$$





Techniques based on a construction called as radical inverse

n	Binary	$\Phi_b(n)$
1	1	0.1
2	01	0.01
3	11	0.11
4	001	0.001
5	101	0.101

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Radical inverse:

 $\Phi_b(n) = 0.d_1d_2...d_m$





Techniques based on a construction called as radical inverse

n	Binary	$\Phi_b(n)$
1	1	0.1 = 1/2
2	01	0.01 = 1/4
3	11	0.11 = 3/4
4	001	0.001 = 1/
5	101	0.101 = 5/



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Radical inverse:

 $\Phi_b(n) = 0.d_1d_2...d_m$







Techniques based on a construction called as radical inverse

n	Binary	$\Phi_b(n)$
1	1	0.1 = 1/2
2	01	0.01 = 1/4
3	11	0.11 = 3/4
4	001	0.001 = 1/
5	101	0.101 = 5/



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Halton and Hammerslay Sequence

Radical inverse: $\Phi_b(n) = 0.d_1d_2...d_m$



Techniques based on a construction called as radical inverse

- Halton Sequence: For n-dimensional sequence, we use different base b for each dimension
 - $x_i = (\Phi_2(i), \Phi_3(i), \Phi_5(i), \dots, \Phi_{p_n}(i))$


Halton and Hammerslay Sequence

Radical inverse: $\Phi_b(n) = 0.d_1d_2...d_m$

$$x_i = (\Phi_2(i), \Phi_3(i), \Phi_5(i), \dots, \Phi_{p_n}(i))$$

Hammerslay Sequence: All except the first dimension has co-prime bases

$$x_i = \left(\frac{i}{N}, \Phi_{b_1}(i), \Phi_{b_2}(i), \dots, \Phi_{b_n}(i)\right)$$



Techniques based on a construction called as radical inverse

Halton Sequence: For n-dimensional sequence, we use different base b for each dimension



Halton and Hammerslay Sequence

Techniques based on a construction called as radical inverse

Radical inverse: $\Phi_b(n) = 0.d_1d_2...d_m$

Halton Sequence:

$$x_i = (\Phi_2(i), \Phi_3(i), \Phi_5(i), \dots, \Phi_{p_n}(i))$$

Hammerslay has slightly lower discrepancy than Halton

Hammerslay Sequence:

$$x_i = \left(\frac{i}{N}, \Phi_{b_1}(i), \Phi_{b_2}(i), \dots, \Phi_{b_n}(i)\right)$$

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Visualizing samples



Figure 2.7: Hammersley Point Set on the 2D Plane. Three 2-dimensional Hammersley point sets $\mathbf{P}_{HAM}^2 = \left(\frac{i}{N}, \Phi_2(i)\right)_{i \in (0,...,N-1)}$ of sizes N = 64-element, N = 256-element and N = 512-element.









Visualizing samples



Figure 2.5: Halton sequence. The first 64, 256, and 512 points of the 2-dimensional Halton Sequence $\mathbf{P}_{HAL}^2 = (\Phi_2(i), \Phi_3(i))_{i \in \mathbb{N}_0}$.









Visualizing samples

Projection: (19,20) Projection: (9,10)





Projection: (29,30)

Halton Sequence

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Faure's permutation



Figure 2.12: Halton Sequence and Scrambled Halton Sequence, Dimensions 7 and 8. (a) The first 256 elements of the 2-dimensional Halton sequence $P_{HAL}^2 =$ $(\Phi_7(i), \Phi_8(i))$ and the scrambled versions of dimension 7 and 8 generated according to procedure of Faure.





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Quasi-Monte Carlo Integration

• Low discrepancy sequences



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Quasi-Monte Carlo Integration

- Low discrepancy sequences
 - Van der Corpus, Sobol sequences



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Quasi-Monte Carlo Integration

- Low discrepancy sequences
 - Van der Corpus, Sobol sequences
 - (t,m,s)-nets & (t-s)-sequences



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Discrepancy



Figure 2.2: Star Discrepancy and Extreme Discrepancy. Visualization of the discrepancy concepts—case s=2—introduced in Definition 2.2. The star discrepancy based on axis-aligned 2-dimensional subareas of I^2 attached at the origin, and the extreme discrepancy based on the choice of arbitrary 2-dimensional subvolumes of \mathbf{I}^2 ...









Discrepancy

of \mathbf{P} is defined as

 $D_N(\mathbf{P}) \equiv D_N$

def sup $\mathbf{B} \in \mathfrak{P}$

where \mathfrak{B} corresponds to a Lebesgue measurable family of subsets of \mathbf{I}^{s} , # corresponds to the counting measure over \mathcal{B} with respect to P, μ^{s} is, as usual, the Lebesgue measure and \mathbf{B} refers to a non empty subset of B.



DEFINITION 2.1 (Discrepancy) Let $\mathbf{P} = \{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N\}$ with $\mathbf{x}_i \in \mathbf{I}^s, i = 1, \dots, N$ be a point set. The discrepancy of P, denoted as $D_N(P)$, is a measure for the deviation of a point set from its ideal distribution. The discrepancy

$$\mathbf{P}_{B} \left| \frac{\#(\mathbf{P} \cap \mathbf{B})}{N} - \mu^{s}(\mathbf{B}) \right|,$$







Fourier Analysis: Quality Measure

Advance Sampling Strategies: June 7, 2018



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